

Introduction to Time Series Analysis of Macroeconomic- and Financial-Data

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Lecture 2: Testing & Dependence over time

Yesterday:

- Econometric Models
- Ordinary Least-Squares (OLS) Regression
 - Line of best fit
 - Hypothesis testing
 - Goodness of fit
 - More than one explanatory variable

Today:

- Is the model well-specified?
- Time-series: dependence over time!

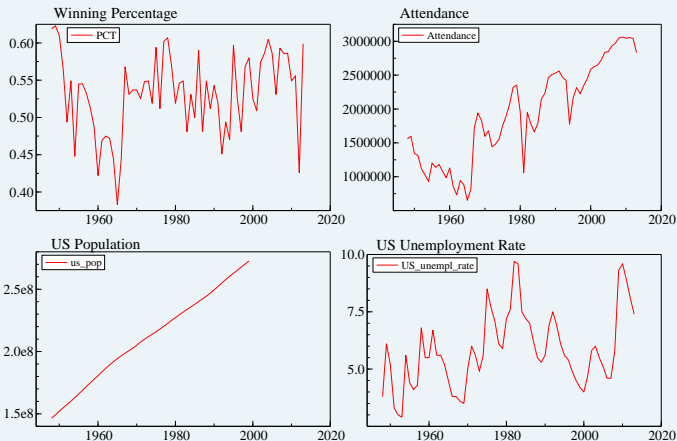
- Extend model to include **two explanatory variables**: X_t and Z_t .
- Triplets of data, (Y_t, X_t, Z_t) : Three economic variables occurring simultaneously.
- Assumptions:
 - (i) (Y_t, X_t, Z_t) **independent** across t .
 - (ii) Identical conditional distribution:
 $(Y_t|X_t, Z_t) \sim (\beta_1 + \beta_2 X_t + \beta_3 Z_t, \sigma^2)$.
 - (iii) X_t and Z_t **exogenously** determined for Y_t .
 - (iv) A parameter space exists: $\beta_1, \beta_2, \beta_3, \sigma^2 \in \mathbb{R}^{\neq} \times \mathbb{R}^+$,
- Gives model:

$$Y_t = \beta_1 + \beta_2 X_t + \beta_3 Z_t + \epsilon_t, \epsilon_t \sim \mathbb{N}0, \sigma^2 \quad (1)$$

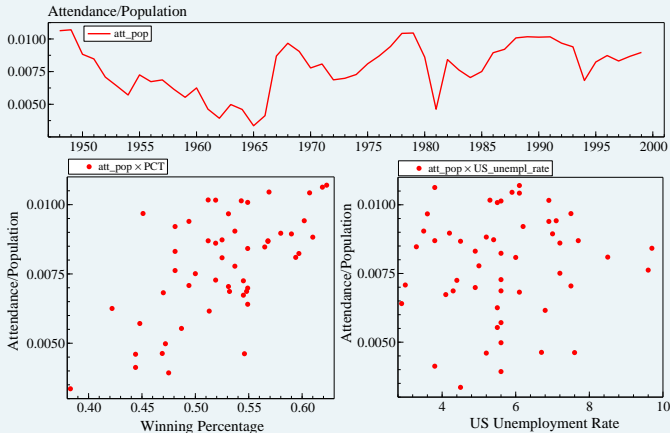
- Parameter interpretation as before except **ceteris paribus**:

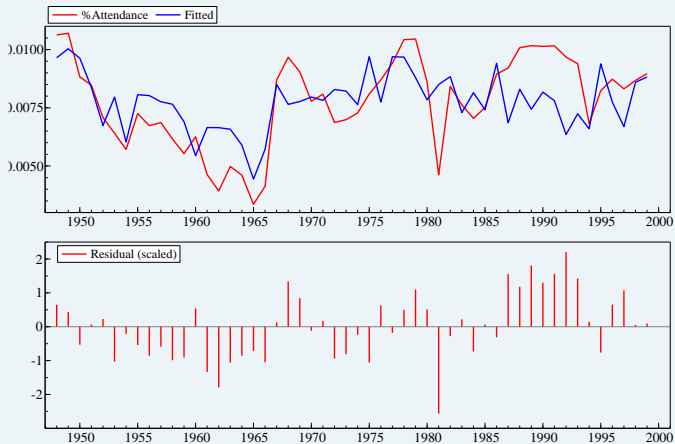
$$\frac{\partial \mathbb{E}Y_t | X_t, Z_t}{\partial X_t} = \beta_2, \quad (2)$$

- Red Sox Game Attendance 1948-1999
- Winning Percentage (PCT)
- US Populations - to scale Attendance (att_pop)
- US Unemployment Rate



$$\text{Model: \%Attendance}_t = \beta_1 + \beta_2 \text{PCT}_t + \beta_3 \text{Unemp.}_t + \epsilon_t$$





EQ(1) Modelling att_pop by OLS

The dataset is: red_sox_1948.in7

The estimation sample is: 1948 - 1999

	Coefficient	Std.Error	t-value	t-prob	Part.R ²
Constant	-0.00475283	0.002263	-2.10	0.0409	0.0826
PCT	0.0224757	0.003977	5.65	0.0000	0.3946
US_unempl_rate	0.000129058	0.0001368	0.943	0.3502	0.0178

sigma	0.00151203	RSS	0.000112024814
R ²	0.398903	F(2,49) =	16.26 [0.000]**
Adj.R ²	0.374369	log-likelihood	265.464
no. of observations	52	no. of parameters	3
mean(att_pop)	0.00779052	se(att_pop)	0.00191161

When the log-likelihood constant is NOT included:

AIC	-12.9326	SC	-12.8201
HQ	-12.8895	FPE	2.41812e-006

When the log-likelihood constant is included:

AIC	-10.0948	SC	-9.98220
HQ	-10.0516	FPE	4.13002e-005

AR 1-2 test:	F(2,47)	=	8.0881 [0.0010]**
ARCH 1-1 test:	F(1,50)	=	1.2062 [0.2773]
Normality test:	Chi ² (2)	=	0.12384 [0.9400]
Hetero test:	F(4,47)	=	1.5738 [0.1969]
Hetero-X test:	F(5,46)	=	1.2577 [0.2983]
RESET23 test:	F(2,47)	=	0.67768 [0.5127]

- Create **null hypothesis**: $H_0 : \beta_2 = 0$.
 - Variable is insignificant: Regression coefficient has zero mean.
- Divide estimator ($\hat{\beta}_2$ here) by estimated standard error ($\sqrt{V[\hat{\beta}_2]}$ here).
 - This calculates the t-statistic.
- Check t-statistic against critical values from t distribution (t_T here):
 - If 5% significance level then check against tables or take roughly ± 2 .¹
- If t-statistic greater than ± 2 then we **reject null hypothesis**.
 - Otherwise we **do not reject null hypothesis**.

¹Recall test is two-tailed: We do not care whether significance is positive or negative!

- In regression output, the t - $prob$ column contains **p-values**.
- In output of other tests (details later), p-values are in [square brackets].
- p-value is probability of rejecting null-hypothesis if null hypothesis is true
- probability of observing test-statistic under the null hypothesis
 - if "sufficiently unlikely" (small), reject H_0
 - Calculated by imposing null hypothesis, using data and assumed distribution.
- If p-value very small, null hypothesis is statistically unlikely:
 - If p-value below significance level α (e.g. 0.05) we **reject null**.
- p-value **probability of falsely rejecting null hypothesis**.

Baseball Model:

$$\text{Att_pop}_t = -0.004753 + 0.02248 \text{PCT}_t + 0.0001291 \text{Unemp}_t$$

(0.00226) (0.00398) (0.000137)

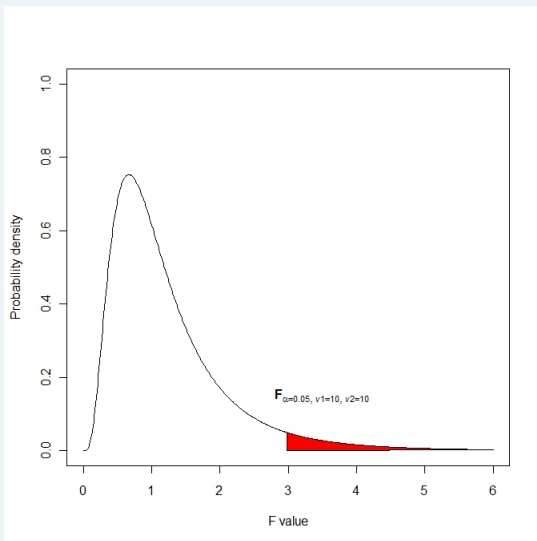
- In multivariate regression, F-test of joint significance different from t-statistics.
- t-testing is for individual regression coefficients:
 - Testing the (null) hypothesis that each one individually is insignificant.
- F-testing is for combinations of regression coefficients:
 - Testing the (null) hypothesis that together they are insignificant.
- Unemployment individually insignificant, variables jointly significant?

$$F(2, 49) = 16.26 [0.000] **$$

- We test joint hypotheses using the F-test:
 - Calculate test statistic and find p-value: Probability of observing test statistic under null
- We find test statistic and distribution:

$$F = \frac{(RSS_R - RSS_U)/(K_U - K_R)}{RSS_U/(T - K_U)} \sim F_{K_U - K_R, T - K_U}. \quad (3)$$

- Baseball model: $H_0 : \beta_2 = 0.025, \beta_3 = 0$.
 - We *could* impose null by creating $Y_t - X_t$ variable but much easier: Use PcGive:
 - *Test then General Restrictions.*
 - Each variable denoted by ampersand (&) and number (See key beneath).
 - Each restriction is line of code; **must** be ended with semi-colon ; .



- Common use of F test: Overall significance of model.
 - More relevant when we add **even more regressors**: Thursday and Friday.

- F-test statistic:

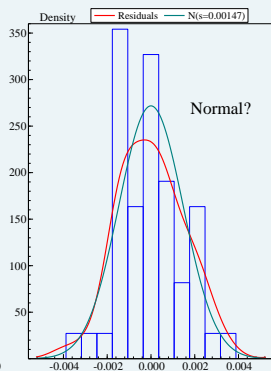
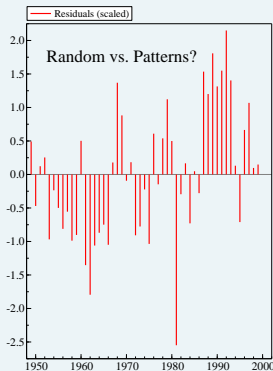
$$F = \frac{(RSS_R - RSS_U)/(K_U - K_R)}{RSS_U/(T - K_U)} = \frac{(TSS - RSS)/(K - 1)}{RSS/(T - K)}. \quad (4)$$

- $ESS = TSS - RSS$: How much do variables explain of variation?
 - Above constant: Base model is still one-variable model
 $Y_t = \beta_1 + \epsilon_t$.
- Form of F yields distribution: $F_{K-1, T-K}$ and hence **critical values**.
 - Null hypothesis: Explanatory variables are **jointly statistically insignificant**.
 - E.g. $H_0 : \beta_2 = \beta_3 = 0$.

- Our assumptions imply ϵ_t has **independent** (i) **identical** (ii) distribution.
 - Often written as $\epsilon_t \sim \text{iid}(0, \sigma^2)$ (distribution need not be Normal).

$$Y_t = \underbrace{\beta_1 + \beta_2 X_t + \beta_3 Z_t}_{\text{Our model: What we know}} + \underbrace{\epsilon_t}_{\text{Errors: What we don't know}}, \quad \epsilon_t \sim \text{iid}(0, \sigma^2). \quad (5)$$

- **Important** because want errors (mistakes) from our model to be:
 - Unrelated to each other.
 - Always of a similar size.
- This is **central principle of all economic and econometric modelling**.
- If not then our model does not include some important information!
 - More importantly, $\hat{\beta}_2, \hat{\beta}_3$ may be **biased** and **inefficient**.



AR 1-2 test:	F (2, 47)	=	8.0881	[0.0010] **
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- Estimated regression model:

$$Y_t = \hat{\beta}_1 + \hat{\beta}_2 X_t + \hat{\beta}_3 Z_t + \hat{\epsilon}_t. \quad (6)$$

- Everything so far depends on iid assumption holding:
 - Unbiasedness and efficiency of estimators.
 - Accuracy of standard errors and all other test statistics.
- Must test whether assumptions hold: What do $\hat{\epsilon}_t$ look like?
- Tests called **misspecification tests** or **diagnostic tests**.

Assume $\epsilon_t \sim \text{iid}(0, \sigma^2) \Rightarrow$

Model:

- $X_{1,t}, \dots, X_{K,t}$.
- Functional form.
- Data transformations.

\iff **Check**

1 Identical distribution:

- Is $\epsilon_t \sim (0, \sigma^2)$?
- Mean must be zero (if constant included), but variance can change.
- **Heteroskedasticity**: $V\epsilon_t = \sigma_t^2$.

2 Independent distribution:

- Is $\text{Corr}[\epsilon_s, \epsilon_t] = E[\epsilon_s \epsilon_t] = 0$, for all $s \neq t$? If not, **autocorrelation**.
- Variance not autocorrelated (**ARCH**)

3 Normal distribution:

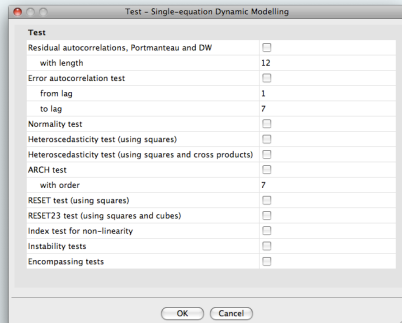
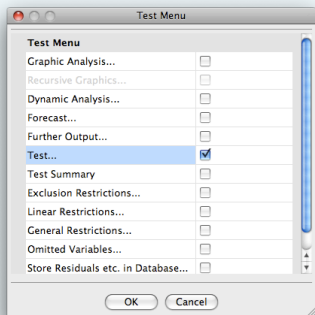
- Is $\epsilon_t \sim N(0, \sigma^2)$?
- Residuals may have very different distribution.

4 Functional Form:

- Model specification test

⇒ will consider in turn!

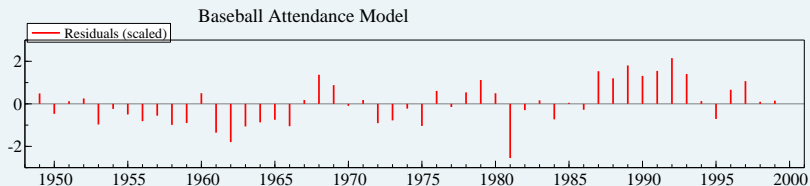
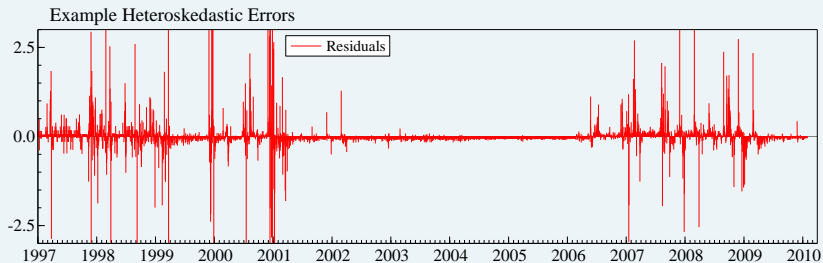
- OxMetrics has very developed ways (tests) for detecting misspecification.



- When conducting testing in OxMetrics, will come across many different test.
 - t-test: Tests on individual regression coefficients.
 - F-test: Tests on combinations of regression coefficients.
- Additionally three types of test based on Likelihood:
 - **Maximum likelihood** is alternative, more flexible form of estimation.
 - Choose parameters (β_1, β_2) to maximise likelihood of having observed our data.
 - Likelihood function is probability density as function of parameters.
- Tests are based on different aspects of likelihood function:
 - **Likelihood Ratio (LR) test**: How likely is null hypothesis?
 - **Wald test**: Computes only under alternative hypothesis.
 - **Lagrange Multiplier (LM) test**: Evaluates only under null hypothesis.
- Type of test not too important; **p-values are the fundamental concept.**

- Assumption of $\epsilon_t \sim (0, \sigma^2)$ violated.
 - Variance changes through sample: $\epsilon_t \sim (0, \sigma_t^2)$.
- Examples:
 - Stock market more volatile in financial crisis.
 - Inflation more volatile when higher.
- Loss of precision:
 - **Efficiency** of estimator relies on constant σ^2 (homoskedasticity).
 - Standard errors larger than they should be.
 - Possible mistake: May conclude variable insignificant that is actually significant.
- On average we still get right answer, but more possibility **we make mistakes**.

- Detection: Use residuals $\hat{\epsilon}_t$ in place of errors ϵ_t .
- Graphic inspection: Does the variance 'look' constant?
 - Residuals (scaled) are standard graphics output in OxMetrics



- White test of heteroskedasticity: regress $\hat{\epsilon}_t^2$ on $X_{1,t}, \dots, X_{K,t}$, and $X_{1,t}^2, \dots, X_{K,t}^2$:
 - Does residual variance depend on regressors included in model, or their squares?
 - Square of variable rough approximation to variance.
- Model:

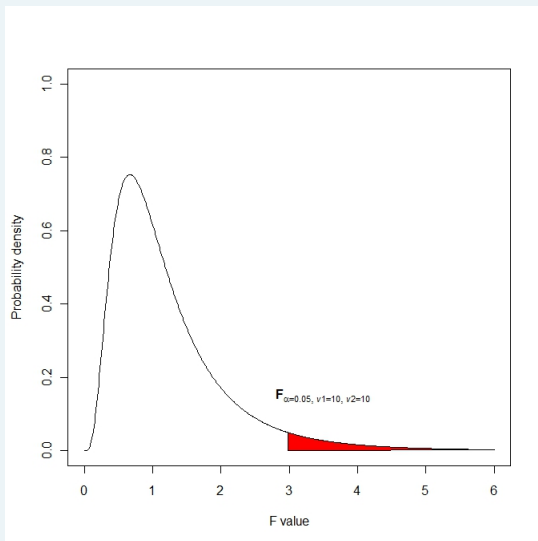
$$\hat{\epsilon}_t^2 = \alpha_0 + \alpha_1 X_{1,t} + \dots + \alpha_K X_{K,t} + \alpha_{K+1} X_{1,t}^2 + \dots + \alpha_{2K} X_{K,t}^2 + v_t.$$

- Null hypothesis:

$$H_0 : \alpha_1 = \dots = \alpha_{2K} = 0, \quad (7)$$

- Absence of heteroskedasticity, or homoskedasticity (constant variance).
- F-test of joint significance:

$$F_{\text{Het}} = \frac{R_{\text{Het}}^2/m}{(1 - R_{\text{Het}}^2)/(T - m)} \sim F_{m, T-m}. \quad (8)$$



Formal Test in Attendance Model:

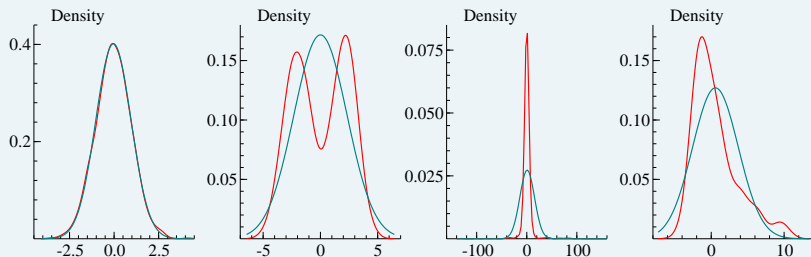
- Hetero test : $F(4, 47) = 1.5738[0.1969]$

Causes:

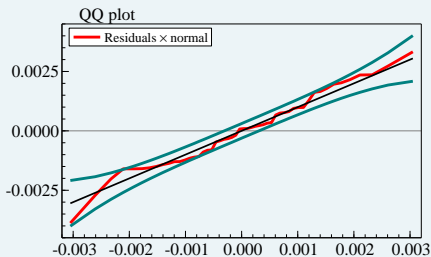
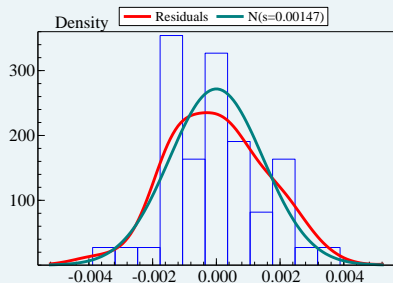
- Omitted Variables:
 - Systematic variation in residuals.
 - Unincluded variable may explain change in variation.
 - E.g. Age may affect how volatile income is across individuals.
- Wrong functional form/data transformation:
 - Logarithmic transformation can stabilise variance in series.
- Non-constant parameters:
 - We are assuming β_1, β_2 stay same throughout sample.
 - If they don't then result may be heteroskedasticity.

- Active:
 - Add the regressors you think cause heteroskedasticity.
- Passive: Calculate Heteroskedasticity Consistent Standard Errors (HCSE).
 - Heteroskedasticity affects standard errors: Become bigger.
 - These standard errors are robust to heteroskedasticity.

- Do not assume Normal distribution but Normal distribution is 'nice' distribution.
- Assume $\epsilon_t \sim \text{iid}(0, \sigma^2)$, so:
 - Distribution unchanging over time.
 - Distribution symmetric around zero.



- Normality not assumed for OLS so provided iid assumption holds inference unaffected:
 - OLS estimator unbiased, consistent, and efficient.
- **But:** Non-Normality may reveal information about dataset:
 - Inappropriate functional form and/or data transformation.
- Informally “test” by plotting data series:



- Skewness and kurtosis of distribution are like mean and variance:
 - They describe 'shape' of distribution:
 - Skewness: How symmetric is the distribution?
 - Kurtosis: How flat/spiky is distribution?
- Test statistics:

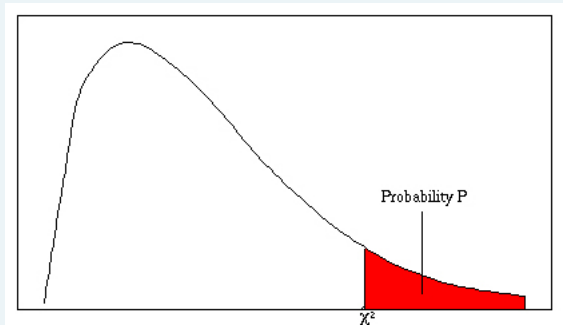
$$\chi^2_{\text{skewness}} = T \frac{\hat{\kappa}_3^2}{6} \sim \chi^2_1, \quad (9)$$

$$\chi^2_{\text{kurtosis}} = T \frac{\hat{\kappa}_4^2}{24} \sim \chi^2_1, \quad (10)$$

$$\chi^2_{\text{normality}} = \chi^2_{\text{skewness}} + \chi^2_{\text{kurtosis}} \sim \chi^2_2. \quad (11)$$

- PcGive reports this test. Null hypothesis is Normal residuals.
 - Hence if test rejected (we see stars) then our model is **misspecified**.
 - Test in Baseball model: Normality test: $\text{Chi}^2(2) = 0.12384[0.9400]$

Under null-hypothesis:



- Example: Normality test: $\text{Chi}^2(2) = 0.12384[0.9400]$
- We should investigate non-Normalities:
 - 1 Non-Normality may be harmless for inference.
 - 2 Non-Normality may reveal important information about model specification.
- 1 Non-Normality may be harmless for inference:
 - iid assumption translates into **non-skewed distribution**.
 - Hence if test fail caused by **excess kurtosis** then model is fine.
- 2 May reveal information about model to help diagnose other problems found.
 - Inappropriate data transformation.
 - Structural breaks (see recursive testing later)
- Additional information via Test menu and `Test . . .`

- Economic theory often predicts which variables matter.
 - Less often predicts mathematical form of dependence.
- Test for functional form: RESET test (Ramsey, 1969).²
 - Include squares, cubes of fitted values.
 - Null hypothesis of correct functional form: Additional variables do not matter.
- Auxiliary regression:

$$Y_i = \beta_1 + \beta_2 X_{2,i} + \beta_3 X_{3,i} + \beta_4 X_{4,i} + \psi_1 \hat{Y}_i^2 + \psi_2 \hat{Y}_i^3 + v_i.$$

- Null hypothesis:

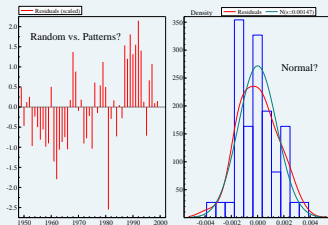
$$H_0 : \psi_1 = \psi_2 = 0. \quad (12)$$

- F-test statistic:

$$F_{\text{RESET}} = \frac{(RSS_R - RSS_U)/2}{RSS_U/(T - K - 2)} \sim F_{2, T-K-2}. \quad (13)$$

- RESET23 test: $F(2, 47) = 0.67768 [0.5127]$

²REGression Specification Error Test.



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Common Model Specification: Log Model

- So far:

$$Y_t = \beta_1 + \beta_2 X_t + \epsilon_t \quad (14)$$

- $\beta_2 = \frac{dY_t}{dX_t}$
- Log-log Model: interpret as percentage changes/elasticities

$$\log(Y)_t = \beta_1 + \beta_2 \log(X_t) + \epsilon_t \quad (15)$$

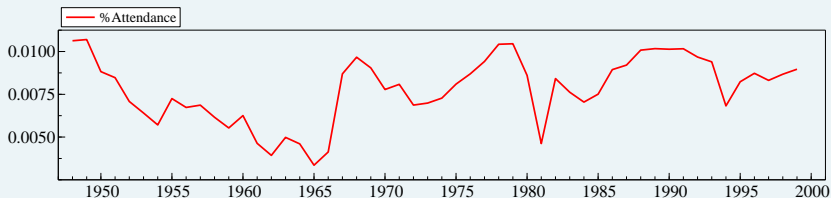
- $\beta_2 = \frac{d \log(Y)_t}{d \log(X)_t} \approx \frac{\frac{dY_t}{Y_t}}{\frac{dX_t}{X_t}} = \frac{dY_t}{dX_t} \frac{X_t}{Y_t} =$ elasticity of Y_t with respect to X_t
- Percentage change! For a 1% change in X_t expect a $\beta_2\%$ change in Y_t .

- So far: Test theory using econometric model.
 - E.g. Baseball attendance: positive relationship between winning percentage and attendance.
- **Model:** $\%Attendance_t = \beta_1 + \beta_2 PCT_t + \beta_3 Unemp._t + \epsilon_t$
- So far: All variables dated t : This is a **static** regression.
 - But where variables are today depends on where they were previously.
 - How can we predict variables if all our dates are t ?
- Today: **Dynamic** models.

- May be that variable is correlated with itself at a previous point in time.
 - E.g.: Exchange rate, interest rates, inflation (low and stable prices), unemployment, baseball attendance.
- Correlation through time is concept called **time dependence**.
 - We recognise how economic variable today is dependent on where it was yesterday.
- Expect persistence in many economic variables; especially prices.
 - Supply reflects production capacity which does not drastically change.
 - Demand reflects frequent purchases if a necessity good.

- Y_{t-r} is **lagged variable**:
 - The same variable at previous point in time.
- Y_{t-1} is first lag: First previous time period.
- Y_{t-2} is second lag, Y_{t-3} third, so on...
- Use lags to learn about:
 - Persistence: How important is a variable's history?
 - Effects of explanatory variables distributed over time (e.g. advertising).

- E.g. Spot exchange rate: Demand and supply of UK pounds for Japanese Yen (Top), % Attendance (bottom).



- Want to understand persistence:
 - Tells us much about economic variables.
 - E.g. price efficiency, partial adjustments, interest rate smoothing.
 - If we don't model it properly, can cause big mistakes.
- **Autoregressive models:**
 - Regression model of variable Y_t on itself in previous time period Y_{t-1} .

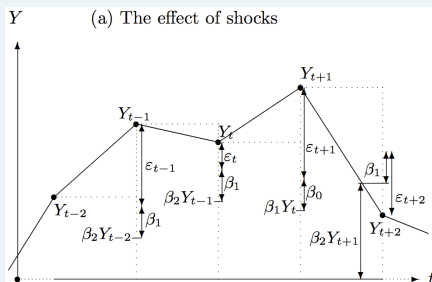
- Autoregressive model has three elements:

- (1) Where Y_t was the last time period.
- (2) The unexpected event ϵ_t .
- (3) Constant term allowing mean of Y_t to be non-zero.

$$Y_t = \underbrace{\alpha_0}_{(3)} + \underbrace{\alpha_1 Y_{t-1}}_{(1)} + \underbrace{\epsilon_t}_{(2)}, \quad \epsilon_t \sim N[0, \sigma^2]. \quad (16)$$

Notation: normally use α for autoregressive, but equivalent to:

$$Y_t = \beta_1 + \beta_2 Y_{t-1} + \epsilon_t, \quad \epsilon_t \sim N[0, \sigma^2]. \quad (17)$$



- AR(1) model allows us to determine many things about theory:
 - α_1 : How quickly equilibrium re-established.
 - α_0 and α_1 : Whether equilibrium is zero or otherwise.
 - σ^2 : How much variation there is in Y_t around equilibrium.
 - How big are the unexpected events?
- What is equilibrium value? Again expectations:

$$EY_t = \alpha_0 + \alpha_1 EY_{t-1}. \quad (18)$$

- Since $EY_t = EY_{t-1}$ we find that $\mu_Y = EY = \alpha_0 / (1 - \alpha_1)$.
 - We define μ_Y to be the equilibrium value, or **unconditional mean** of Y_t .

- We learn about the persistence of deviations from equilibrium from α_1 .
- To see why note that $\mu_Y = \alpha_0 / (1 - \alpha_1)$ implies $\alpha_0 = \mu_Y(1 - \alpha_1)$ so that:

$$Y_t = \alpha_0 + \alpha_1 Y_{t-1} + \epsilon_t \implies Y_t - \mu_Y = \alpha_1 (Y_{t-1} - \mu_Y) + \epsilon_t. \quad (19)$$

- We have **de-meaned** Y_t : We only care about α_1 and deviations from equilibrium.
- If assume **no more shocks happen** can see how quickly impact of shock disappears.
- $Y_t - \mu_Y = \alpha_1 (Y_{t-1} - \mu_Y)$ and $Y_{t-1} - \mu_Y = \alpha_1 (Y_{t-2} - \mu_Y)$ so:

$$Y_t - \mu_Y = \alpha_1^2 (Y_{t-2} - \mu_Y). \quad (20)$$

- We can carry on doing this:

$$Y_t - \mu_Y = \alpha_1^3(Y_{t-3} - \mu_Y) \quad \dots \quad Y_t - \mu_Y = \alpha_1^k(Y_{t-k} - \mu_Y) \quad (21)$$

- It so happens that:

$$\text{Corr}[Y_t, Y_{t-k}] = \frac{\text{Cov}(Y_t, Y_{t-k})}{\sqrt{V(Y_t)}\sqrt{V(Y_{t-k})}} = \frac{\alpha_1^k \sigma_Y^2}{\sigma_Y \times \sigma_Y} = \alpha_1^k. \quad (22)$$

- Have measured autocorrelation, or correlation through time, of Y_t from α_1 !
- The bigger is α_1 and hence nearer to 1, the more persistent is the series:
 - If $\alpha_1 = 0.9$ then $\alpha_1^2 = 0.81$ and $\alpha_1^{10} = 0.35$.
 - If $\alpha_1 = 0.2$ then $\alpha_1^2 = 0.04$ and $\alpha_1^{10} \approx 0$.

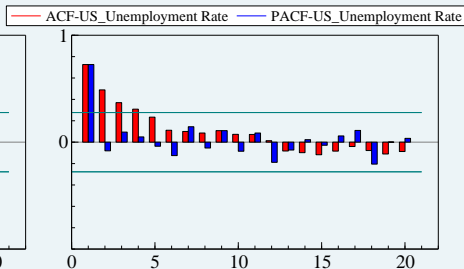
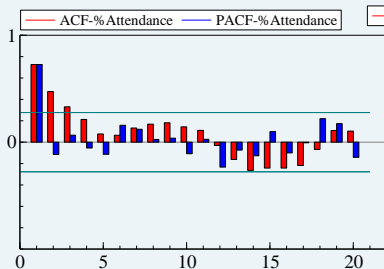
- May need more than one lag to explain dynamics of variable:
 - If we model p lags, we have AR(p) model.
- E.g. AR(2): $Y_t = \alpha_0 + \alpha_1 Y_{t-1} + \alpha_2 Y_{t-2} + \epsilon_t$.
- Estimators like in **multivariate regression**:
 - $\hat{\alpha}_2$ asks Y_{t-1} to **be still!** It controls for first lag to get **only second lag effect**.

$$\hat{\alpha}_2 = \frac{\sum_{t=2}^T Y_{t-2}(Y_t|Y_{t-1})}{\sum_{t=2}^T Y_{t-2}(Y_{t-2}|Y_{t-1})}$$

- Unconditional mean, variance and covariance affected. E.g. unconditional mean:

$$\mu_Y = \frac{\alpha_0}{1 - \alpha_1 - \alpha_2}. \quad (23)$$

- Common method for learning about autocorrelation is graphically.
 - Autocorrelation function (ACF): $\text{Corr}[Y_t, Y_{t-p}]$, $p = 1, 2, \dots, 20$.
 - Partial ACF (PACF): $\text{Corr}[Y_t, Y_{t-p} | Y_{t-1}, \dots, Y_{t-p+1}]$, $p = 1, 2, \dots, 20$.
 - Number of significant PACF lags = number of lags needed in model.



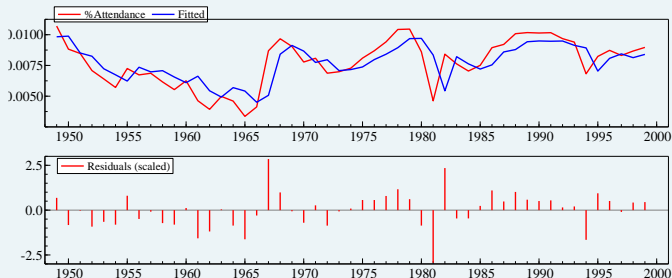
- Autoregressive model for Baseball Attendance:

$$\%Attendance_t = \mu_0 + \alpha_1 \%Attendance_{t-1} + \epsilon_t \quad (24)$$

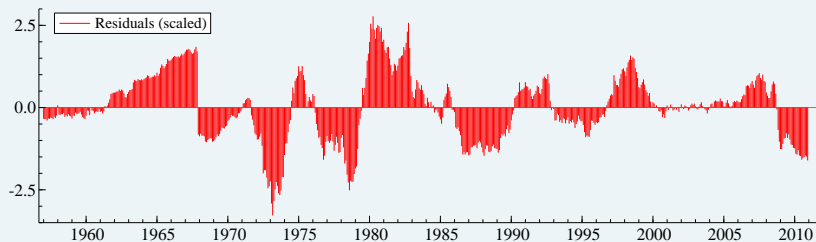
- Estimated Model:

$$\%Attendance_t = 0.002052 + 0.7316 \%Attendance_{t-1}$$

(0.000747) (0.0935)



- Earlier: Misspecification testing: Check iid assumption.
 - $\hat{\epsilon}_t$ all approximately same size?
 - $\hat{\epsilon}_t$ all unrelated to each other?
- Latter failure in time series is **residual autocorrelation**:
 $\text{Corr}(\epsilon_t, \epsilon_{t-1}) \neq 0$.
- Omitted variable bias: We omit lags that matter for explaining Y_t .



- Test using auxiliary regression:

$$\hat{\epsilon}_t = \gamma_0 + \gamma_1 Y_{t-1} + \phi_1 \hat{\epsilon}_{t-1} + v_t, \quad (25)$$

- Null hypothesis:

$$H_0 : \gamma_1 = \phi_1 = 0, \quad (26)$$

- Absence of autocorrelation in errors of model (model well specified).
- F-test statistic:

$$F_{AR} = \frac{R_{AR}^2 / r}{(1 - R_{AR}^2) / (T - r)} \sim F_{r, T-r}. \quad (27)$$

- Can include up to r lags when testing. Can vary in Test menu and
- Remedy autocorrelation by including time-dependent information:
 - More lags or extra variables (with lags).

- ARCH: **A**utoregressive **C**onditional **H**eteroskedasticity.
- Heteroskedasticity: $\epsilon_t \sim (0, \sigma_t^2)$, variance changes over time.
- Autoregressive: Variable correlated with itself.
- ARCH: $\sigma_t^2 = \alpha_0 + \alpha_1 \epsilon_{t-1}^2 + u_t$.
 - Variance is persistent: Spells of high and low volatility.
 - Common in financial markets.
- Impact on regression model same as heteroskedasticity:
Efficiency.
- Test using AR(1) on squared residuals:
 $\hat{\epsilon}_t^2 = \alpha_0 + \alpha_1 \hat{\epsilon}_{t-1}^2 + \dots + \alpha_r \hat{\epsilon}_{t-r}^2 + v_t$.
 - F-test: $H_0 : \alpha_1 = \dots = \alpha_r = 0$.

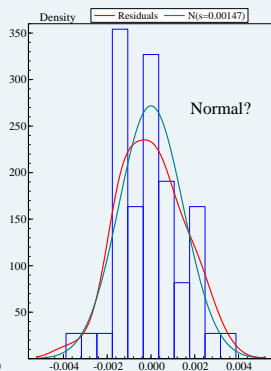
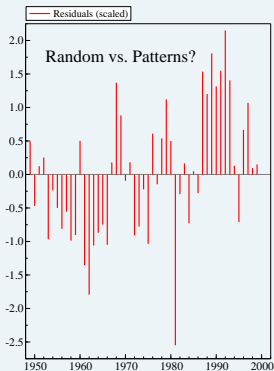
- Recall our model: **Model:**

$$\%Attendance_t = \beta_1 + \beta_2 PCT_t + \beta_3 Unemp._t + \epsilon_t$$

- Misspecification Tests

AR 1-2 test:	F (2, 47)	=	8.0881	[0.0010]**
ARCH 1-1 test:	F (1, 50)	=	1.2062	[0.2773]
Normality test:	Chi ² (2)	=	0.12384	[0.9400]
Hetero test:	F (4, 47)	=	1.5738	[0.1969]
Hetero-X test:	F (5, 46)	=	1.2577	[0.2983]
RESET23 test:	F (2, 47)	=	0.67768	[0.5127]

→ Residual Autocorrelation!



AR 1-2 test:	F (2, 47)	=	8.0881	[0.0010] **
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RESET23 test:	F (2, 47)	=	0.67768	[0.5127]

Add extra lags of dependent variable:

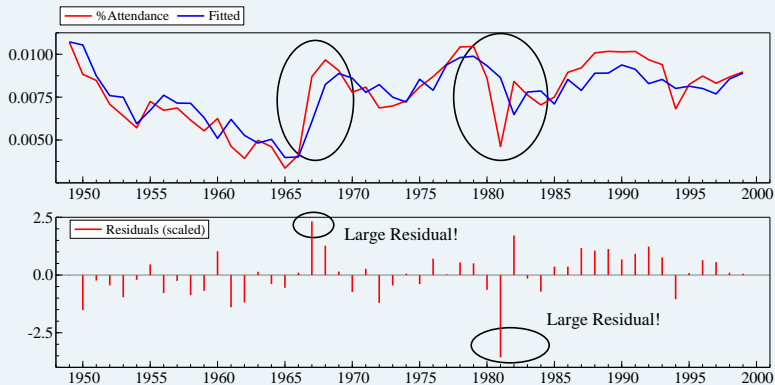
$$\%Attend._t = \alpha_1 \%Attend._{t-1} + \beta_1 + \beta_2 PCT_t + \beta_3 Unemp._t + \epsilon_t$$

$$\begin{aligned} \%Attend. &= 0.5815 \%Attend._{t-1} - 0.003985 + 0.01307 PCT_t \\ &\quad (0.0913) \qquad\qquad\qquad (0.00171) \qquad\qquad\qquad (0.00335) \\ &+ 0.00006 Unempl._t \\ &\quad\qquad\qquad (0.000104) \end{aligned}$$

AR 1-2 test:	F (2, 45)	=	0.60392	[0.5510]
ARCH 1-1 test:	F (1, 49)	=	0.67357	[0.4158]
Normality test:	Chi ² (2)	=	11.261	[0.0036] **
Hetero test:	F (6, 44)	=	0.53186	[0.7810]
Hetero-X test:	F (9, 41)	=	0.43033	[0.9109]
RESET23 test:	F (2, 45)	=	0.10631	[0.8994]

What now??

Look at the model fit and residuals:



May distort normality tests! What happened? More on Thursday!

Computer Lab Session 2:

- Testing joint-hypotheses
 - F-test
- Interpreting Misspecification Tests
- Investigating Time Series Properties
 - Time Series plots

Fulton Fish Market: Price, Quantity, Weather

- Load "fish.in7"
- Series
 - Model for $q_t = \log(\text{Quantity})$
 - Weather: Stormy, Rainy, Cold
- Graph the series! (Important first step!)

Construct Auto-regressive models for $\log(\text{Quantity})$ sold:

- Determine lag length: (Partial) Auto-correlation function (max 10 lags)
- Estimate an AR(1), AR(2) models
- What is the long-run equilibrium?
- Interpret mis-specification tests

EQ(15) Modelling qty by OLS

The estimation sample is: 2 - 111

	Coefficient	Std.Error	t-value	t-prob	Part.R ²
qty_1	0.203549	0.09406	2.16	0.0327	0.0416
Constant	6.78385	0.8049	8.43	0.0000	0.3968

sigma	0.731432	RSS		57.7792493	
R ²	0.0415598	F(1,108) =	4.683	[0.033]*	
Adj.R ²	0.0326853	log-likelihood		-120.671	
no. of observations	110	no. of parameters		2	
mean(qty)	8.51915	se(qty)		0.743687	

AR 1-2 test:	F(2,106)	=	1.9872	[0.1422]
ARCH 1-1 test:	F(1,108)	=	2.0874	[0.1514]
Normality test:	Chi ² (2)	=	6.9103	[0.0316]*
Hetero test:	F(2,107)	=	3.6890	[0.0282]*
Hetero-X test:	F(2,107)	=	3.6890	[0.0282]*
RESET23 test:	F(2,106)	=	0.69995	[0.4989]

Interested in **effects of weather** on quantity sold:

- Estimate auto-regressive model with weather variables added in
- Include: Stormy, Rainy, Cold
- Which variables are individually significant?
- Which variables are jointly significant?

EQ(17) Modelling qty by OLS

	Coefficient	Std.Error	t-value	t-prob	Part.R ²
qty_1	0.184254	0.09336	1.97	0.0511	0.0358
Constant	7.06086	0.8097	8.72	0.0000	0.4200
stormy	-0.342175	0.1681	-2.04	0.0443	0.0380
rainy	0.0824118	0.1918	0.430	0.6683	0.0018
cold	-0.0566163	0.1524	-0.372	0.7109	0.0013

sigma	0.721793	RSS	54.7034867
R ²	0.0925804	F(4,105) =	2.678 [0.036]*
Adj.R ²	0.0580121	log-likelihood	-117.663
no. of observations	110	no. of parameters	5
mean(qty)	8.51915	se(qty)	0.743687

AR 1-2 test:	F(2,103) =	0.82520 [0.4410]
ARCH 1-1 test:	F(1,108) =	1.7838 [0.1845]
Normality test:	Chi ² (2) =	8.7179 [0.0128]*
Hetero test:	F(5,104) =	1.3869 [0.2352]
Hetero-X test:	F(5,104) =	1.3869 [0.2352]
RESET23 test:	F(2,103) =	0.65843 [0.5198]

F-Test for joint significance of all 3 weather variables

- *Test - Exclusion Restrictions*
- Select all 3 variables

Test for excluding:

[0] = stormy

[1] = rainy

[2] = cold

Subset F(3, 105) = 1.9679 [0.1233]

Modelling share price (Apple) using Google search data:

Can share price be modelled using consumer interest (measured through google search term: "iphone")

- load data "apple_google.in7"
- Build a model for Apple closing price
- What lag length should you choose?
- Is the google search term a useful predictor? What about lagged search terms?
- What about the effect of google searches on the trading volume?